

ON DISTRIBUTION FUNCTION OF GENERALIZED ORDER STATISTICS FROM THE LOMAX-WEIBULL DISTRIBUTION

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Abstract- The Lomax-Weibull distribution [11] was defined by combining the cumulative distribution functions of well-known Lomax and Weibull distributions. In this paper, we derive the joint distribution function of two generalized order statistics (GOS) when the independent random variables are from continuous Lomax-Weibull distribution. Then the joint distribution of two order statistics from Lomax-Weibull distribution is derived as a special case of the main result.

Keywords: Lomax-Weibull distribution, joint distribution, generalized order statistics, cumulative distribution.

MSC Subject Classification: 60E, 62N.

1. INTRODUCTION

Order statistics finds many important applications in statistics and statistical modelling of various processes such as robust location estimates, detection of outliers, censored sampling and reliability to mention a few. These models describe random variables arranged in increasing order of magnitude.

If the random variables $X_1, X_2, \dots X_n$ are arranged in increasing order of magnitudes and then written as $X_{(1)} \leq X_{(2)} \leq \dots \leq X_{(n)}$, then $X_{(i)}$ is called the ith order statistics where $i = 1, \dots, n$. The unordered random variables X_i are generally statistically independent and having identical distribution but the ordered random variables X(i), $i = 1, 2, \dots, n$ are need to be dependent.

Generalized order statistics is of interest since past few years because they are more suitable in statistical modelling of various real life problems, reliability theory and statistical inference. This concept was introduced as a unified distribution system to study various models of ordered random variables with different representations or interpretations. This concept of order statistics was further studied and extended to the concept of GOS. Kamps in his papers and books [16-18] did a pioneer work in this area. He proved how GOS generalizes the concepts of order statistics, record values, and other ordered random variables. This concept was further studied by various researchers Ahsanullah [1-3], El-Baset, Ahmed and Al-Matrofi [4], Cramer and Kamps [5-6], Cramer, Kamps and Rychlik [7-9], Kamps and Cramer [14] and Garg [10] to mention a few.

The Weibull distribution is extensively used in the modelling of data in reliability, biological studies and also in analysing life time data while the Lomax distribution is very important in modelling business problems, economics and actuarial modelling due to its heavy tail, therefore in this paper we have used the Lomax-Weibull distribution [11] which is more flexible and generalized to both Lomax and Weibull distributions.

In this paper we shall derive the joint distribution function of two GOS from the Lomax-Weibull distribution. Then the joint distribution function of order statistics from Lomax-Weibull distribution is derived as the special case of our main result.

2. PRELIMINARIES

2.1 The Lomax-Weibull distribution

The pdf of the Lomax-Weibull distribution defined by Gupta and Garg [] is as follows

$$f_{LW}(x;\alpha,\beta,\lambda) = \begin{cases} \frac{k\alpha C}{\beta\lambda} \left(\frac{x}{\lambda}\right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k} \left(1 + \frac{e^{-\left(\frac{x}{\lambda}\right)^k}}{\beta}\right)^{-(1+\alpha)}, & x \ge 0\\ 0, & x < 0 \end{cases}$$
(1)

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And the cdf by

$$F_{LW}(x;\alpha,\beta,\lambda) = \begin{cases} 1-C\left\{1-\left(1+\beta^{-1}e^{-\left(\frac{x}{\lambda}\right)^{k}}\right)^{-\alpha}\right\} & , & x \ge 0\\ 0 & , & x < 0 \end{cases}$$
(2)
Where

$$C^{-1} = 1 - \left(1 + \beta^{-1}\right)^{-\alpha}, \qquad (3)$$

$$\beta, \lambda, k > 0 \text{ and } \alpha \neq 0.$$

2.2 Generalized Order Statistics

Let us denote the cumulative density function of an absolutely continuous function with F(x) and density function with f(x) and $X_{i;n,m,k}$ and $X_{j;n,m,k}$ ($k \ge 1$, m is a real number) denote the ith and jth GOS derived from a random sample of size 'n' then the joint pdf $f_{i,j;n,m,k}(x_i, x_j)$ of ith and jth generalized order statistics is given by

$$f_{i,j;n,m,k}(x_i, x_j) = \frac{C_j}{(i-1)!(j-i-1)!} \Big[1 - F(x_i) \Big]^m \Big[1 - F(x_j) \Big]^{\gamma_j - 1} \Big[g_m(F(x_i)) \Big]^{i-1} \Big]$$

$$\left[g_m(F(x_j)) - g_m(F(x_i)) \Big]^{j-i-1} f(x_i) f(x_j) \right]$$
(4)

 $0 \le x_i < x_j < \infty$, $1 \le i < j \le n$

$$C_{j} = \prod_{r=1}^{j} \gamma_{r}, \qquad \gamma_{r} = k + (n-r)(m+1)$$

$$g_{m}(x) = \begin{cases} \frac{1}{m+1} \left[1 - (1-x)^{(m+1)} \right], & m \neq -1 \\ -\ln(1-x) &, m = -1, x \in (0,1) \end{cases}$$
(5)
(6)

Since

$$\lim_{m \to -1} \frac{1}{m+1} \left[1 - (1-x)^{(m+1)} \right] = -\ln(1-x)$$

$$g_m(x) = \frac{1}{m+1} \left[1 - (1-x)^{(m+1)} \right] \text{ for all } x \in (0,1) \text{ and for all } m$$

With

 $g_{-1}(x) = \lim_{m \to -1} g_m(x)$

3. THE JOINT DISTRIBUTION FUNCTION OF TWO GOS FROM LOMAX-WEIBULL DISTRIBUTION

Theorem Let $X_{i;n,m,k}$ and $X_{j;n,m,k}$ denote the ith and jth GOS from a random sample of size 'n' drawn from Lomax-Weibull distribution [1-2], then the joint pdf of the ith and jth GOS is given by

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$$\begin{cases} \frac{C_{j} \alpha^{2} \lambda^{2} K^{1+m+\gamma_{j}+(m+1)(j-i-1)}}{\beta^{2} (i-1)! (j-i-1)!} \left(\frac{1}{m+1}\right)^{j-2} \sum_{l_{i}=0}^{j-1} \sum_{l_{2}=0}^{j-i-1} (-1)^{l_{i}+l_{2}} K^{l_{i}(m+1)} \binom{i-1}{l_{1}} \binom{j-i-1}{l_{2}} \left(1+\frac{e^{-\lambda x_{i}}}{\beta}\right)^{-(\alpha+1)} \left(1+\frac{e^{-\lambda x_{j}}}{\beta}\right)^{-(\alpha+1)} \\ \cdot e^{-\lambda (x_{i}+x_{j})} \left\{1 - \left(1+\frac{e^{-\lambda x_{i}}}{\beta}\right)^{-\alpha}\right\}^{m+(m+1)(l_{i}-l_{2}+j-i-1)} \left\{1 - \left(1+\frac{e^{-\lambda x_{j}}}{\beta}\right)^{-\alpha}\right\}^{-\gamma_{j}-1+(m+1)l_{2}}; \text{ for } m \neq -1 \end{cases}$$

$$f_{i,j;n,m,k} (x_{i}, x_{j}) = \begin{cases} \frac{k^{j} \alpha^{2} \lambda^{2} K^{k^{j}} e^{-\lambda (x_{i}+x_{j})}}{\beta^{2} (i-1)! (j-i-1)!} \left\{1 - \left(1+\frac{e^{-\lambda x_{i}}}{\beta}\right)^{-\alpha}\right\}^{-1} \left\{1 - \left(1+\frac{e^{-\lambda x_{j}}}{\beta}\right)^{-\alpha}\right\}^{-1} \left\{1 - \left(1+\frac{e^{-\lambda x_{j}}}{\beta}\right)^{-\alpha}\right\}^$$

Provided that $\alpha, \beta, \lambda > 0, 1 \le i < j \le n, k \ge 1$, m is a real number and n>1, $x_j > x_i \ge 0$ and C_j and γ_j are defined by Equation (40), and $K^{-1} = 1 - (1 + \beta^{-1})^{-\alpha}$.

Proof. The result can easily be established on substituting the values of f(x), F(x), and $g_m(x)$ from Equations (1), (2) and (6) respectively in the Equation (4), expressing

the values of $\left[g_m(F(x_i))\right]^{i-1}$ and $\left[g_m(F(x_j)) - g_m(F(x_i))\right]^{j-i-1}$ in their series forms and doing some simplification.

Special Case. If we take m=0 and k = 1 in the above Theorem, then GOS reduce into order statistics and we get the joint distribution the ith and jth order statistics X_{i,n} and X_{j,n} from a sample of size n from Lomax-Weibull distribution, which is given by

$$f_{i,j;n,k}(x_{i},x_{j}) = \frac{\alpha^{2} \lambda^{2} n! K^{(n+1-i)}}{\beta^{2} (n-j)! (i-1)! (j-i-1)!} \sum_{l_{1}=0}^{i-1} \sum_{l_{2}=0}^{j-i-1} (-1)^{l_{1}+l_{2}} K^{l_{1}} \binom{i-1}{l_{1}} \binom{j-i-1}{l_{2}} \left(1 + \frac{e^{-\lambda x_{i}}}{\beta}\right)^{-(\alpha+1)} \left(1 + \frac{e^{-\lambda x_{j}}}{\beta}\right)^{-(\alpha+1)} \\ \cdot e^{-\lambda (x_{i}+x_{j})} \left\{1 - \left(1 + \frac{e^{-\lambda x_{i}}}{\beta}\right)^{-\alpha}\right\}^{(l_{1}-l_{2}+j-i-1)} \left\{1 - \left(1 + \frac{e^{-\lambda x_{j}}}{\beta}\right)^{-\alpha}\right\}^{-(\alpha+1)}$$

$$(8)$$

4. ESTIMATION

We estimate the parameters of Lomax-Weibull distribution by the method of maximum likelihood. The log-likelihood for a random sample $x_1, ..., x_n$ from pdf (1) is

$$\log L(\alpha, \beta, \lambda, k) = \log(k\alpha) - n\log(\lambda) - n\log(\beta) + (k-1)\sum_{i=1}^{n}\log\frac{x_i}{\lambda} - k\sum_{i=1}^{n}\frac{x_i}{\lambda} - n\log\left\{1 - \left(1 + \frac{1}{\beta}\right)^{-\alpha}\right\} - (\alpha+1)\sum_{i=1}^{n}\log\left[1 + \frac{1}{\beta}\exp\left(-\frac{x_i}{\lambda}\right)^{k}\right]$$
(9)

The first order partial derivatives of Log L, Equation (9), with respect to the three parameters are

$$\frac{\partial \log L}{\partial \alpha} = \frac{1}{\alpha} - \frac{n\left(1+\beta^{-1}\right)^{-\alpha}\log\left(1+\beta^{-1}\right)}{1-\left(1+\beta^{-1}\right)^{-\alpha}} - \sum_{i=1}^{n}\log\left[1+\frac{1}{\beta}\exp\left(-\frac{x_{i}}{\lambda}\right)^{k}\right]$$
(10)

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$$\frac{\partial \log L}{\partial \beta} = -\frac{n}{\beta} + \frac{\alpha n \left(1 + \beta^{-1}\right)^{-1}}{\beta^2 \left[\left(1 + \beta^{-1}\right)^{\alpha} - 1 \right]} + \frac{(\alpha + 1)}{\beta^2} \sum_{i=1}^n \frac{\exp\left(-\frac{x_i}{\lambda}\right)^k}{\left[1 + \beta^{-1} \exp\left(-\frac{x_i}{\lambda}\right)^k\right]}$$
(11)

$$\frac{\partial \log L}{\partial \lambda} = -\frac{n}{\lambda} - \frac{n(k-1)}{\lambda} + \sum_{i=1}^{n} \frac{x_i}{\lambda^2} - \frac{(\alpha+1)}{\beta \lambda^2} \sum_{i=1}^{n} \frac{k x_i \left(\frac{x_i}{\lambda}\right) - \exp\left(-\frac{x_i}{\lambda}\right)}{\left[1 + \beta^{-1} \exp\left(-\frac{x_i}{\lambda}\right)^k\right]}$$
(12)

$$\frac{\partial \log L}{\partial k} = \frac{1}{k} + \sum_{i=1}^{n} \log \frac{x_i}{\lambda} - \sum_{i=1}^{n} \frac{x_i}{\lambda} - \frac{(\alpha+1)}{\beta} \sum_{i=1}^{n} \frac{\left(\frac{x_i}{\lambda}\right)^k \log\left(\frac{x_i}{\lambda}\right)^k \exp\left(-\frac{x_i}{\lambda}\right)^k}{\left[1 + \beta^{-1} \exp\left(-\frac{x_i}{\lambda}\right)^k\right]}$$
(13)

Setting these expressions to zero and solving them simultaneously yields the maximum-likelihood estimates of the four parameters.

CONCLUSION

Looking into the wide application, the joint distribution function of two generalized order statistics when the independent random variables are from continuous Lomax-Weibull distribution is derived. Also the joint distribution of two order statistics from Lomax-Weibull distribution is derived as a special case of the main result.

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